A universal error bound in the CLT for counting monochromatic edges in uniformly colored graphs

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Abstract

Let \( \{G_n : n \geq 1\} \) be a sequence of simple graphs. Suppose \( G_n \) has \( m_n \) edges and each vertex of \( G_n \) is colored independently and uniformly at random with \( c_n \) colors. Recently, Bhattacharya, Diaconis and Mukherjee (2014) proved universal limit theorems for the number of monochromatic edges in \( G_n \). Their proof was by the method of moments, and therefore was not able to produce rates of convergence. By a non-trivial application of Stein’s method, we prove that there exists a universal error bound for their central limit theorem. The error bound depends only on \( m_n \) and \( c_n \), regardless of the graph structure.

Keywords: Stein’s method; normal approximation; rate of convergence; monochromatic edges.

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1 Introduction

Let \( \{G_n : n \geq 1\} \) be a sequence of simple graphs, that is, graphs that contain no loops and no multiple edges. Suppose \( G_n \) has \( m_n \) edges and each vertex of \( G_n \) is colored independently and uniformly at random with \( c_n \) colors. Let \( Y_n \) be the number of monochromatic edges in \( G_n \). Using the coupling approach in Stein’s method for Poisson approximation, Barbour, Holst and Janson [3] (page 105, Theorem 5.G) proved that

\[
d_{TV}(\mathcal{L}(Y_n), \text{Poi}(\frac{m_n}{c_n})) \leq \frac{\sqrt{8m_n}}{c_n}
\]

(1.1)

where \( d_{TV} \) denotes the total variation distance and \( \text{Poi}(\lambda) \) denotes the Poisson distribution with mean \( \lambda \). The bound (1.1) implies that if \( c_n \to \infty \) and \( m_n/c_n \to \lambda > 0 \), the distribution of \( Y_n \) converges to the Poisson distribution with mean \( \lambda \). Recently, Bhattacharya, Diaconis and Mukherjee [4] reproved this Poisson limit theorem by the method of moments. By the same method, they also showed that in the case \( c_n \to \infty \) and \( m_n/c_n \to \infty \), the distribution of \( Y_n \), after proper standardization, converges to the standard normal distribution. These limit theorems were called universal limit theorems because they do not require any assumption on the graph structure. For applications of this and related problems, we refer to [4] and the references therein.

In this note, we prove the following result.

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Theorem 1.1. Let $Y$ be the number of monochromatic edges in a simple graph with $m \geq 1$ edges where each vertex is colored independently and uniformly at random with $c \geq 2$ colors. Let

$$W = \frac{(Y - \frac{m}{c})}{\sqrt{\frac{m}{c}(1 - \frac{1}{c})}}.$$ 

We have

$$d_W(\mathcal{L}(W), N(0,1)) \leq 3 \sqrt{\frac{\sqrt{c}}{m}} + \frac{10\sqrt{2}}{\sqrt{c}} + \frac{1}{\sqrt{\pi}} \frac{2^{7/4}}{m^{1/4}}. \quad (1.2)$$

where $d_W$ denotes the Wasserstein distance and $N(0,1)$ denotes the standard normal distribution.

The bound (1.2) provides a universal error bound for the central limit theorem for $W$ as $c \to \infty$ and $m/c \to \infty$. A corollary for fixed $c$ is also obtained in Remark 2.5. The bound (1.2) is obtained by a non-trivial application of Stein’s method for normal approximation.

Stein’s method was introduced by Stein [15] for normal approximation. Stein’s method for Poisson approximation was first studied by Chen [7] and popularized by Arratia, Goldstein and Gordon [1]. We refer to [2] for an introduction to Stein’s method. Stein’s method has been widely used to prove limit theorems with error bounds in graph theory. For example, Arratia, Goldstein and Gordon [1] and Chatterjee, Diaconis and Meckes [6] used Stein’s method to prove Poisson limit theorems for monochromatic cliques in a uniformly colored complete graph. Cerquetti and Fortini [5] considered more general monochromatic subgraphs counts when the distribution of colors was exchangeable. Janson and Nowicki [12] studied the asymptotic distribution of the number of copies of a given graph in various random graph models.

All of the above results are obtained by exploiting the local dependence structure within random variables. Chen and Shao [8] provides general normal approximation results for sums of locally dependent random variables. Rinott and Rotar [14] gives multivariate normal approximation results for sums of bounded locally dependent random vectors. Their error bounds are on the Kolmogorov distance and on the difference between probabilities on multidimensional convex sets. However, these results are not directly applicable in our problem, which does not have a useful (LD2) structure required in these two papers. In addition to the local dependence structure, we also exploit the uncorrelatedness within $W$. This technique of exploiting the uncorrelatedness within random variables was also used in [10] to obtain rates of convergence for the central limit theorem for subgraph counts in random graphs.

We would like to mention that Goldstein and Rinott [11] studied the more general multidimensional version of the graph coloring problem. Moreover, their result does not require the number of colors to go to infinity or the coloring to be uniform. However, they only considered regular graphs with degree of each vertex not growing too fast. For example, their bound does not converge to 0 for complete graphs. We only consider the case where the number of colors goes to infinity since it is a necessary condition to ensure that the central limit theorem holds for all graphs with sufficiently large number of edges, regardless of the graph structure and therefore universal.

In the next section, we give the proof of Theorem 1.1.

2 Normal approximation

Let $G = (V(G), E(G))$ be a simple undirected graph, where $V(G)$ is the vertex set and $E(G)$ is the edge set. Let $m = |E(G)|$ be the number of edges of $G$. We color each vertex of $G$ independently and uniformly at random with $c \geq 2$ colors. Formally, label
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the vertices of $G$ by $\{v_1, \ldots, v_{|V(G)|}\}$ and denote the color of the vertex $v_i$ by $\xi_v$. Label the edges of $G$ by $\{1, \ldots, m\}$. For each edge $i$, we denote by $v_{i1}$, $v_{i2}$ the two vertices it connects, i.e., $i = (v_{i1}, v_{i2})$. Without loss of generality, assume $\deg(v_{i1}) \leq \deg(v_{i2})$ where $\deg(v)$ denotes the degree of vertex $v$. Using the above notation, the standardized number of monochromatic edges can be expressed as

$$W = \sum_{i=1}^{m} X_i := \sum_{i=1}^{m} \left( \frac{I(\xi_{v_{i1}} = \xi_{v_{i2}}) - \frac{1}{2}}{\sqrt{\frac{m}{e}(1 - \frac{1}{2})}} \right).$$ (2.1)

Observing that $X_i$ and $X_j$ are uncorrelated if $i \neq j$, we have $\mathbb{E}W = 0$, $\text{Var}(W) = 1$.

We will need the following lemmas in the proof of Theorem 1.1.

**Lemma 2.1.** We have the following bounds on the moments of $X_i$:

$$\mathbb{E}|X_i| \leq \frac{2}{\sqrt{m} e}; \quad \mathbb{E}X_i^2 = \frac{1}{m}; \quad \mathbb{E}|X_i|^3 \leq \frac{\sqrt{e}}{m^{3/2}}.$$ (2.2)

**Proof.** The proof is elementary and therefore omitted. \qed

**Lemma 2.2** (Page 37 of [3]). For each edge $i = (v_{i1}, v_{i2})$, define $d_i = \deg(v_{i1}) \land \deg(v_{i2})$. We have

$$\sum_{i=1}^{m} d_i \leq \sqrt{2} m^{3/2}.$$ (2.3)

**Lemma 2.3** (Lemma 2.2 of [4]). The number of triangles, denoted by $\#(\Delta)$, in $G$ is bounded by $\sqrt{2} m^{3/2}/3$.

The following proposition is the key ingredient in proving Theorem 1.1.

**Proposition 2.4.** For any function $f$ with bounded first and second derivatives, we have with $W$ defined in (2.1),

$$\|\mathbb{E}f'(W) - \mathbb{E}f(W)\| \leq \|f''\| \left( \frac{3}{2} \sqrt{\frac{e}{m}} + \frac{5\sqrt{2}}{\sqrt{e}} \right) + \|f'\| \frac{2^{1/4}}{m^{1/4}}$$ (2.4)

where $\|g\| := \sup_x |g(x)|$ for any function $g$.

**Proof.** For each edge $i = (v_{i1}, v_{i2})$ with $\deg(v_{i1}) \leq \deg(v_{i2})$, define the neighborhood $N_i \subset \{1, \ldots, m\}$ to consist of all edges connect to $v_{i1}$. Let

$$D_i = \sum_{j \in N_i} X_j, \quad W_i = W - D_i.$$

Since the color of $v_{i1}$ is independent of $W_i$, we have $X_i$ is independent of $W_i$. Let $U$ be an independent random variable distributed uniformly in $[0, 1]$. By $\mathbb{E}X_i = 0$, $\mathbb{E}X_i^2 = 1/m$, the Taylor expansion and adding and subtracting corresponding terms, we have

$$\mathbb{E}f'(W) - \mathbb{E}f(W) = \sum_{i=1}^{m} \mathbb{E}X_i^2 \mathbb{E}f'(W) - \sum_{i=1}^{m} \mathbb{E}X_i[f(W) - f(W_i)]$$

$$= \sum_{i=1}^{m} \mathbb{E}X_i^2 \mathbb{E}f'(W) - \sum_{i=1}^{m} \mathbb{E}X_i D_i f'(W - UD_i)$$

$$= \sum_{i=1}^{m} \mathbb{E}X_i^2 \mathbb{E}f'(W) - \sum_{i=1}^{m} \mathbb{E}X_i^2 f'(W - UD_i) - \sum_{i=1}^{m} \mathbb{E}X_i (D_i - X_i) f'(W - UD_i)$$

$$=: R_1 - R_2 - R_3 - R_4$$ (2.5)
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where

\[ R_1 = \sum_{i=1}^{m} \mathbb{E}X_i^2\mathbb{E}f'(W) - \sum_{i=1}^{m} \mathbb{E}X_i^2\mathbb{E}f'(W_i), \]
\[ R_2 = \sum_{i=1}^{m} \mathbb{E}X_i^2[f'(W - UD_i) - f'(W_i)], \]
\[ R_3 = \sum_{i=1}^{m} \mathbb{E}X_i(D_i - X_i)[f'(W - UD_i) - f'(W)], \]
\[ R_4 = \mathbb{E}f'(W)\sum_{i=1}^{m} X_i(D_i - X_i). \]

First of all, by the Taylor expansion,

\[ |R_1| \leq ||f''|| \sum_{i=1}^{m} \mathbb{E}X_i^3\mathbb{E}|D_i| \leq ||f''|| \left( \sum_{i=1}^{m} \mathbb{E}|X_i|^3 + \sum_{i=1}^{m} \sum_{j \in N_i \setminus \{i\}} \mathbb{E}|X_i|^2\mathbb{E}|X_j| \right). \]

By (2.2) and (2.3),

\[ \sum_{i=1}^{m} \mathbb{E}|X_i|^3 \leq \sqrt{\frac{c}{m}} \sum_{i=1}^{m} \sum_{j \in N_i \setminus \{i\}} \mathbb{E}|X_i|^2\mathbb{E}|X_j| \leq \frac{1}{m} \frac{2}{mc} \sum_{i=1}^{m} d_i \leq \frac{2\sqrt{3}}{\sqrt{c}}. \]

Therefore,

\[ |R_1| \leq ||f''||\left( \sqrt{\frac{c}{m}} + \frac{2\sqrt{3}}{\sqrt{c}} \right). \]

By the same argument and the fact that \( \{X_j : j \in N_i\} \) are jointly independent,

\[ |R_2| \leq \frac{1}{2} ||f''||\left( \sqrt{\frac{c}{m}} + \frac{2\sqrt{2}}{\sqrt{c}} \right). \]

For \( R_3 \), by the Taylor expansion,

\[ |R_3| \leq \frac{1}{2} ||f''|| \sum_{i=1}^{m} \mathbb{E}|X_i||D_i - X_i||D_i| \leq \frac{1}{2} ||f''|| \sum_{i=1}^{m} \mathbb{E}|X_i|^2|D_i - X_i| + \frac{1}{2} ||f''|| \sum_{i=1}^{m} \mathbb{E}|X_i||D_i - X_i|^2. \]

Again by (2.2), (2.3) and the fact that \( \{X_j : j \in N_i\} \) are jointly independent,

\[ \sum_{i=1}^{m} \mathbb{E}|X_i|^2|D_i - X_i| \leq \frac{1}{m} \frac{2}{mc} \sum_{i=1}^{m} d_i \leq \frac{2\sqrt{2}}{\sqrt{c}}, \]
\[ \sum_{i=1}^{m} \mathbb{E}|X_i||D_i - X_i|^2 \leq \frac{2}{mc} \sum_{i=1}^{m} d_i \leq \frac{2\sqrt{2}}{\sqrt{c}}. \]

Therefore,

\[ |R_3| \leq ||f''||\frac{2\sqrt{2}}{\sqrt{c}}. \]

Finally we bound \( |R_4| \). By the Cauchy-Schwartz inequality, the fact that \( \{X_j : j \in N_i\} \) are jointly independent and \( \mathbb{E}X_i = 0 \),

\[ |R_4| \leq ||f'|| \sqrt{\text{Var}(\sum_{i=1}^{m} X_i(D_i - X_i))} = ||f'|| \sqrt{\text{Var}(\sum_{j \in N_i \setminus \{i\}} X_iX_j)}. \]
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Observe that if \( j \in N_i \setminus \{i\} \) and \( l \in N_k \setminus \{k\} \), \( \text{Cov}(X_i X_j, X_k X_l) = 0 \) unless \( \{i, j\} = \{k, l\} \) or \( \{i, j, k, l\} \) forms a triangle. For the case \( \{i, j\} = \{k, l\} \),

\[
\text{Cov}(X_i X_j, X_i X_j) = \frac{1}{m^2}.
\]

For the case \( \{i, j, k, l\} \) forms a triangle, with distinct \( i, j, k \),

\[
\text{Cov}(X_i X_j, X_j X_k) = \mathbb{E} X_i X_j^2 X_k \leq \frac{1}{m^2}
\]

where the last inequality is by straightforward calculation. Therefore, by (2.3), Lemma 2.3, and observing that each triangle in \( G \) gives rise to 3 ordered pairs of \( \{i, j\} \) such that \( j \in N_i \setminus \{i\} \), we have,

\[
|R_4| \leq ||f'|| \left\lfloor \sum_{i=1}^{m} \sum_{j \in N(i)} \frac{1}{m^2} + \frac{3 \times 2}{m^2} \#(\Delta) \leq ||f'|| \cdot \frac{2^{1/4}}{m^{1/4}}. \right\rfloor
\]

The bound (2.4) follows from (2.5) and the bounds on \( |R_1| - |R_4| \).

**Proof of Theorem 1.1.** By the definition of the Wasserstein distance, we have

\[
d_W(\mathbb{L}(W), N(0, 1)) = \sup_{||h'|| \leq 1} |\mathbb{E} h(W) - \mathbb{E} h(Z)|.
\]

where \( Z \) is a standard Gaussian random variable. Let \( f_h \) be the solution to

\[
f'(w) - w f(w) = h(w) - \mathbb{E} h(Z)
\]

(2.6)

given by

\[
f_h(w) = e^{w^2/2} \left\lfloor \int_{-\infty}^{w} \{h(t) - \mathbb{E} h(Z)\} e^{-t^2/2} dt. \right\rfloor
\]

Replacing \( w \) by \( W \) and taking expectations on both sides of (2.6), we have

\[
d_W(\mathbb{L}(W), N(0, 1)) = \sup_{||h'|| \leq 1} |\mathbb{E} f_h'(W) - \mathbb{E} W f_h(W)|. \tag{2.7}
\]

If \( ||h'|| \leq 1 \), then it is known that (c.f. (2.14) of [13] and (2.13) of [9])

\[
||f'_h|| \leq \sqrt{\frac{2}{\pi}}, \quad ||f''_h|| \leq 2.
\]

The bound (1.2) is proved by (2.7) and applying the above bounds in (2.4).

**Remark 2.5.** The following bound can be obtained following the proof of Theorem 1.1:

\[
d_W(\mathbb{L}(W), N(0, 1)) \leq C_0 \left( \sqrt{\frac{c}{m}} + \frac{K_m}{\sqrt{m^{3/2}}} + \frac{1}{m^{1/4}} \right)
\]

where \( C_0 \) is an absolute constant, \( c \geq 2 \), \( K_m = \sum_{i=1}^{m} d_i \), and \( d_i \) is defined in Lemma 2.2. For fixed \( c \geq 2 \), the above error bound goes to zero if \( m \to \infty \) and \( K_m = o(m^{3/2}) \). This rules out complete graphs. Theorem 1.3 of [4] shows that a sufficient and necessary condition for the asymptotic normality with fixed \( c \geq 2 \) is \( m \to \infty \) and \( N(C_4) = o(m^2) \) where \( N(C_4) \) is the number of 4-cycles.
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References


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